

Lab8A : IV and RD in Stata

Introduction to Econometrics, Spring 2023

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Section 1

IV estimator in Stata

Subsection 1

Review the Theory

IV estimator in Stata

- Review the Theory

Review Previous Lecture of Internal Validity

Threatens to Internal Validity

- Three endogenous in OLS regression are:
 - Omitted Variable Bias** (a variable that is correlated with X but is unobserved)
 - Simultaneity or reverse causality Bias** (X causes Y , Y causes X)
 - Errors-in-Variables Bias** (X is measured with error)
- One easy way to deal with these endogeneity is using **Instrumental Variable** method.

IV estimator in Stata

- Review the Theory

Instrumental Variable Method

Instrumental variables: 1 endogenous regressor & 1 instrument

- suppose a simple OLS regression like previous equation

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$

- Because $E[u_i|X_i] \neq 0$, then we can use an instrumental variable (Z_i) to obtain an consistent estimate of coefficient.
- Intuitively, we want to split X_i into two parts:
 - part that is correlated with the error term.
 - part that is uncorrelated with the error term.
- If we can isolate the variation in X_i that is uncorrelated with u_i , then we can use this part to obtain a consistent estimate of the causal effect of X_i on Y_i .

IV estimator in Stata

- Review the Theory

Instrumental Variable Method

Instrumental variables: 1 endogenous regressor & 1 instrument

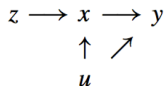
- An instrumental variable Z_i must satisfy the following 2 properties:

- Instrumental relevance:** Z_i should be **correlated** with the casual variable of interest, X_i (endogenous variable), thus

$$\text{Cov}(X_i, Z_i) \neq 0$$

- Instrumental exogeneity:** Z_i is as good as randomly assigned and Z_i only affect on Y_i through X_i affecting Y_i channel.

$$\text{Cov}(Z_i, u_i) = 0$$



- Review the Theory

IV estimator: Two Steps Least Square (2SLS)

- We can estimate the causal effect of X_i on Y_i in two steps

- First stage:** Regress X_i on Z_i & obtain predicted values of \hat{X}_i , if $Cov(Z_i, u_i) = 0$, then \hat{X}_i contains variation in X_i that is uncorrelated with u_i

$$\hat{X}_i = \hat{\pi}_0 + \hat{\pi}_1 Z_i$$

- Second stage:** Regress Y_i on \hat{X}_i to obtain the Two Stage Least Squares estimator $\hat{\beta}_{2SLS}$

$$\hat{\beta}_{2SLS} = \frac{\sum(Y_i - \bar{Y})(\hat{X}_i - \bar{\hat{X}})}{\sum(\hat{X}_i - \bar{\hat{X}})^2}$$

- Review the Theory

IV with Heterogeneous Causal Effects

Instrument Variables: Constant-effect

- Instrumental Variable is a useful method to make causal inference. It can eliminate
 - Omitted Variable Bias
 - Measurement Error
 - Reverse Causality
- Two Assumptions
 - Relevance(Weak Instrument): It can be test by the first stage regression and F-statistic.
 - Exogeneity: Can't be test formally but argue it using professional knowledges.
- Estimation and Inference
 - When IV satisfy these two assumptions,the causal effect of coefficients of interest,TSLS estimator, β_{TSLS} can be NOT unbiased but **consistent**.
 - The sampling distribution of the TSLS estimator is also normal in large samples,so the general procedures for statistical inference in OLS can be used.

IV estimator in Stata

• Review the Theory

Some Practical Guides by Angrist and Pischke(2012)

Practical Guides

1 Check IV relevance

- Always report the first stage and think about whether it makes sense(Signs and magnitudes)
- Always report the F-statistic on the excluded instruments. The bigger,the better. Don't forget the rule of thumb.($F > 10$)

2 Check exclusion restriction

- The exclusion restriction cannot be tested directly, but it can be falsified
- Run and examine the reduced form(regression of dependent variable on instruments) and look at the coefficients, t-statistics and F-statistics for excluded instruments.
- Because the reduced form is proportional to the casual effect of interest and is unbiased(OLS), so we should see the causal relation in the reduced form.If you can't see the causal relation in the reduced form,it's probably not there

• Review the Theory

Some Practical Guides by Angrist and Pischke(2012)

Practical Guides

- 3 Provide a substantive explanation for observed difference between 2SLS and OLS
 - How big is the difference? What does this tell you?
 - Is the coefficient bigger when theory of endogeneity suggests it should be smaller? If so, why?
 - Measurement Error or heterogeneous effects?
- 4 If you have multiple instruments, report over-identification tests.
 - Pick your best single instrument and report just-identified estimates using this one only because just-identified IV is relatively unlikely to be subject to a weak bias.
 - Worry if it is substantially different from what you get using multiple instruments.
 - Check over-identified 2SLS estimates with LIML. LIML is less than precise than 2SIS but also less biased. If the results come out similar, be happy. If not, worry, and try to find stronger instruments.

Subsection 2

Introduction

- Syntax

```
ivregress estimator depvar [varlist1] (varlist2 = varlist_iv)  
^I[if] [in] [weight] [, options]
```

```
help ivregress
```

- Options

- ▶ `estimator` : `2sls/liml/gmm`
- ▶ `devar` : dependent variable
- ▶ `varlist1` : exogenous variables
- ▶ `varlist2` : endogenous variables
- ▶ `varlist_iv` : instrument variables

- Options

- ▶ e.g. $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \varepsilon$,
in which, x_1 is exogenous variables,
 x_2 is endogenous variables,
and z_1, z_2 are instrument variables.

```
ivregress 2sls y x1 (x2 = z1 z2)
ivregress 2sls y x1 (x2 = z1 z2), r first

/* r--异方差稳健标准误
   first--report first-stage regression */
```

Subsection 3

An Example

- An Example

- ▶ Mincer(1958) first researched the positive correlation between **salary** and **years of education**, but omitted the **ability** variable.
- ▶ Griliches(1976) addressed the problem of omitted variable bias with IV method.
- ▶ Data : Young Men's Cohort of the National Longitudinal Survey (NLS-Y).
- ▶ Blackburn and Neumark(1992) Updated data of Griliches(1976).
- ▶ Two-period panel data : the initial period is the earliest year in which the above variables have data; The end period is 1980.

● An Example : Data

- ▶ **lw** (工资对数)
- ▶ **s** (受教育年限)
- ▶ **age** (年龄)
- ▶ **expr** (工龄)
- ▶ **tenure** (在现单位的工作年数)
- ▶ **iq** (智商)
- ▶ **med** (母亲的受教育年限)
- ▶ **kww** (在“knowledge of the World of Work”测试中的成绩)
- ▶ **mrt** (=1, 已婚)
- ▶ **rns** (=1, 住在美国南方)
- ▶ **smsa** (=1, 住在大城市)
- ▶ **year** (有数据的最早年份, 1966-1973 中的某一年)

IV estimator in Stata

1 Data Summary.

```
. use grilic, clear
. sum
```

Variable	Obs	Mean	Std. Dev.	Min	Max
rns	758	.2691293	.4438001	0	1
rns80	758	.292876	.4553825	0	1
mrt	758	.5145119	.5001194	0	1
mrt80	758	.8984169	.3022988	0	1
smsa	758	.7044855	.456575	0	1
smsa80	758	.7124011	.452942	0	1
med	758	10.91029	2.74112	0	18
iq	758	103.8562	13.61867	54	145
kw	758	36.57388	7.302247	12	56
year	758	69.03166	2.631794	66	73
age	758	21.83509	2.981756	16	30
age80	758	33.01187	3.085504	28	38
s	758	13.40501	2.231828	9	18
s80	758	13.70712	2.214693	9	18
expr	758	1.735429	2.105542	0	11.444
expr80	758	11.39426	4.210745	.692	22.045
tenure	758	1.831135	1.67363	0	10
tenure80	758	7.362797	5.05024	0	22
lw	758	5.686739	.4289494	4.605	7.051
lw80	758	6.826555	.4099268	4.749	8.032

- ② Test the correlation between **iq**(智商) and **s**(受教育年限).

```
. pwcorr iq s, star(.01)
```

	iq	s
iq	1.0000	
s	0.5131*	1.0000

```
* Significant positive correlation at the level of 1%,  
* correlation coefficient = 0.51.
```

IV estimator in Stata

- 3 Run an OLS regression.
- With robust standard error
- control variables : **expr tenure rns smsa**
- We are interested in **s**.

```
. reg lw s expr tenure rns smsa, r
```

```
Linear regression
```

```
Number of obs   =       758  
F(5, 752)       =       84.05  
Prob > F        =       0.0000  
R-squared       =       0.3521  
Root MSE       =       .34641
```

lw	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
s	.102643	.0062099	16.53	0.000	.0904523	.1148338
expr	.0381189	.0066144	5.76	0.000	.025134	.0511038
tenure	.0356146	.0079988	4.45	0.000	.0199118	.0513173
rns	-.0840797	.029533	-2.85	0.005	-.1420566	-.0261029
smsa	.1396666	.028056	4.98	0.000	.0845893	.194744
_cons	4.103675	.0876665	46.81	0.000	3.931575	4.275775

- ③ Run an OLS regression.
 - Annual return on investment in education : 10.26% (Significant at the 1% level).
 - **Overestimation** coefficient : omitted the **ability** variable.
 - Ability is positively correlated with years of education.
 - The contribution of ability to wages is included into the contribution of education.

IV estimator in Stata

- Using **iq**(智商) as a **proxy variable** for ability, run an OLS regression.
- Other **proxy variables**: High school test scores; Armed Forces Qualification Test(美国参军资格考试), etc.

```
. reg lw s iq expr tenure rns smsa, r
Linear regression                               Number of obs   =       758
                                                F(6, 751)       =       71.89
                                                Prob > F        =       0.0000
                                                R-squared       =       0.3600
                                                Root MSE       =       .34454
```

lw	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
s	.0927874	.0069763	13.30	0.000	.0790921	.1064826
iq	.0032792	.0011321	2.90	0.004	.0010567	.0055016
expr	.0393443	.0066603	5.91	0.000	.0262692	.0524193
tenure	.034209	.0078957	4.33	0.000	.0187088	.0497092
rns	-.0745325	.0299772	-2.49	0.013	-.1333815	-.0156834
smsa	.1367369	.0277712	4.92	0.000	.0822186	.1912553
_cons	3.895172	.1159286	33.60	0.000	3.667589	4.122754

- ④ Using **iq**(智商) as a **proxy variable** for ability, run an OLS regression.
 - Annual return on investment in education reduced to 9.28%.
 - More reasonable, but still too large.
 - So **iq**(智商) is an **endogenous variable**.

- ⑤ Run 2SLS regression.
 - Consider using **med** (母亲的受教育年限), **kww** (在 “knowledge of the World of Work” 测试中的成绩), **mrt** (=1, 已婚), **age** (年龄) as **IVs** of **iq**(智商).
 - With robust standard error

IV estimator in Stata

5 Run 2SLS regression.

```
Instrumental variables (2SLS) regression
```

Number of obs	=	758
Wald chi2(6)	=	355.73
Prob > chi2	=	0.0000
R-squared	=	0.2002
Root MSE	=	.38336

lw	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
iq	-.0115468	.0056376	-2.05	0.041	-.0225962	-.0004974
s	.1373477	.0174989	7.85	0.000	.1030506	.1716449
expr	.0338041	.0074844	4.52	0.000	.019135	.0484732
tenure	.040564	.0095848	4.23	0.000	.0217781	.05935
rns	-.1176984	.0359582	-3.27	0.001	-.1881751	-.0472216
smsa	.149983	.0322276	4.65	0.000	.0868182	.2131479
_cons	4.837875	.3799432	12.73	0.000	4.0932	5.58255

```
Instrumented: iq  
Instruments: s expr tenure rns smsa med kww mrt age
```

- 5 Run 2SLS regression.
 - Annual return on investment in education increased to 13.73%? (incredible)
 - The contribution of **iq**(智商) to wages is negative? (incredible)
 - We should check **instrument validity**.

IV estimator in Stata

- ⑥ Overidentification test.
 - the number of instruments(4) > the number of endogenous regressors(1)
 - To test instrument exogeneity, thus overidentification test.

```
. estat overid
   Test of overidentifying restrictions:
   Score chi2(3)          = 51.5449 (p = 0.0000)
```

- Compute J-Statistic, some (or one) of the instrumental variables are invalid.
- We suspect `mrt`(=1, 已婚), `age`(年龄) are invalid.
- Compute C-Statistic(检验部分工具变量不满足外生性) using `-ivreg2-`.

```
ssc install ivreg2
```

IV estimator in Stata

6 Overidentification test.

```
/* ivreg2默认估计量为2SLS
   orthog(mrt age):检验(mrt,age)是否满足外生性 */

. ivreg2 lw s expr tenure rns smsa (iq = med kww mrt age), r orthog (mrt age)
IV (2SLS) estimation

Estimates efficient for homoskedasticity only
Statistics robust to heteroskedasticity
```

Total (centered) SS	=	139.2861498	Number of obs =	758
Total (uncentered) SS	=	24652.24662	F(6, 751) =	58.74
Residual SS	=	111.39959	Prob > F =	0.0000
			Centered R2 =	0.2002
			Uncentered R2 =	0.9955
			Root MSE =	.3834

lw	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
iq	-.0115468	.0056376	-2.05	0.041	-.0225962	-.0004974
s	.1373477	.0174989	7.85	0.000	.1030506	.1716449
expr	.0338041	.0074844	4.52	0.000	.019135	.0484732
tenure	.040564	.0095848	4.23	0.000	.0217781	.05935
rns	-.1176984	.0359582	-3.27	0.001	-.1881751	-.0472216
smsa	.149983	.0322276	4.65	0.000	.0868182	.2131479
_cons	4.837875	.3799432	12.73	0.000	4.0932	5.58255

IV estimator in Stata

6 Overidentification test.

```
. ivreg2 lw s expr tenure rns smsa (iq = med kww mrt age), r orthog (mrt age)
Underidentification test (Kleibergen-Paap rk LM statistic):          33.294
                                                                Chi-sq(4) P-val =    0.0000
-----
Weak identification test (Cragg-Donald Wald F statistic):          10.538
                                                                (Kleibergen-Paap rk Wald F statistic):    9.585
Stock-Yogo weak ID test critical values:  5% maximal IV relative bias  16.85
                                           10% maximal IV relative bias  10.27
                                           20% maximal IV relative bias   6.71
                                           30% maximal IV relative bias   5.34
                                           10% maximal IV size           24.58
                                           15% maximal IV size           13.96
                                           20% maximal IV size           10.26
                                           25% maximal IV size            8.31
Source: Stock-Yogo (2005).  Reproduced by permission.
NB: Critical values are for Cragg-Donald F statistic and i.i.d. errors.
-----
Hansen J statistic (overidentification test of all instruments):    51.545
                                                                Chi-sq(3) P-val =    0.0000
-orthog- option:
Hansen J statistic (eqn. excluding suspect orthog. conditions):    0.116
                                                                Chi-sq(1) P-val =    0.7333
C statistic (exogeneity/orthogonality of suspect instruments):    51.429
                                                                Chi-sq(2) P-val =    0.0000
Instruments tested:  mrt age
-----
Instrumented:          iq
Included instruments: s expr tenure rns smsa
Excluded instruments: med kww mrt age
```

⑥ Overidentification test.

- 使用 `-ivreg2-` 得到的回归系数和稳健标准误与 `-ivregress-` 相同;
- 拒绝 (`mrt age`) 满足外生性的原假设;
- 考虑仅使用 (`med kww`) 作为 `iq` 的工具变量。

IV estimator in Stata

7 Run 2SLS regression again

```
. ivregress 2sls lw s expr tenure rns smsa (iq=med kww), r first
First-stage regressions
```

```
Number of obs      =          758
F(   7,   750)     =          47.74
Prob > F           =          0.0000
R-squared          =          0.3066
Adj R-squared      =          0.3001
Root MSE          =         11.3931
```

iq	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
s	2.467021	.2327755	10.60	0.000	2.010052	2.92399
expr	-.4501353	.2391647	-1.88	0.060	-.9196471	.0193766
tenure	.2059531	.269562	0.76	0.445	-.3232327	.7351388
rns	-2.689831	.8921335	-3.02	0.003	-4.441207	-.938455
smsa	.2627416	.9465309	0.28	0.781	-1.595424	2.120907
med	.3470133	.1681356	2.06	0.039	.0169409	.6770857
kww	.3081811	.0646794	4.76	0.000	.1812068	.4351553
_cons	56.67122	3.076955	18.42	0.000	50.63075	62.71169

IV estimator in Stata

7 Run 2SLS regression again

```
Instrumental variables (2SLS) regression
```

Number of obs	=	758
Wald chi2(6)	=	370.04
Prob > chi2	=	0.0000
R-squared	=	0.2775
Root MSE	=	.36436

lw	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
iq	.0139284	.0060393	2.31	0.021	.0020916	.0257653
s	.0607803	.0189505	3.21	0.001	.023638	.0979227
expr	.0433237	.0074118	5.85	0.000	.0287968	.0578505
tenure	.0296442	.008317	3.56	0.000	.0133432	.0459452
rns	-.0435271	.0344779	-1.26	0.207	-.1111026	.0240483
smsa	.1272224	.0297414	4.28	0.000	.0689303	.1855146
_cons	3.218043	.3983683	8.08	0.000	2.437256	3.998831

```
Instrumented: iq  
Instruments: s expr tenure rns smsa med kw
```


- 7 Run 2SLS regression again
 - Annual return on investment in education reduced to 6.08%, which is reasonable.
 - The contribution of `iq`(智商) to wages turns to positive again.

IV estimator in Stata

- 8 Check instrument validity
 - Check **IV relevance** : report the first stage.
 - Instrument perform well in the first stage.
 - A more formal test : F-statistic exceeds 10 (13.40), no Weak Instruments.

```
. estat firststage, all forcenonrobust  
First-stage regression summary statistics
```

Variable	R-sq.	Adjusted R-sq.	Partial R-sq.	Robust F(2,750)	Prob > F
iq	0.3066	0.3001	0.0382	13.4028	0.0000

```
Shea's partial R-squared
```

Variable	Shea's Partial R-sq.	Shea's Adj. Partial R-sq.
iq	0.0382	0.0305

```
Minimum eigenvalue statistic = 14.9058
```

```
Critical Values
```

```
Ho: Instruments are weak
```

```
# of endogenous regressors: 1
```

```
# of excluded instruments: 2
```

	5%	10%	20%	30%
2SLS relative bias				
			(not available)	
2SLS Size of nominal 5% Wald test	10%	15%	20%	25%
	19.93	11.59	8.75	7.25

IV estimator in Stata

- 8 Check instrument validity
 - Run and examine the **reduced form**.

```
qui reg lw s expr tenure rns smsa med, r
est store m1

qui reg lw s expr tenure rns smsa kww, r
est store m2

qui reg lw s expr tenure rns smsa med kww, r
est store m3

esttab m1 m2 m3,
    mtitle("reduced form:med" "reducedform:kww" "reducedform:med,kww")
    b(%6.3f) nogap compress
    star(* 0.1 ** 0.05 *** 0.01)
    ar2 order(med kww)
```

IV estimator in Stata

- 8 Check instrument validity
 - Run and examine the **reduced form**.

	(1) reduced_d	(2) reduced_w	(3) reduced_w
med	0.007 (1.55)		0.007 (1.38)
kww		0.004** (2.07)	0.004* (1.96)
s	0.100*** (15.36)	0.097*** (15.01)	0.095*** (14.06)
expr	0.039*** (5.88)	0.037*** (5.42)	0.037*** (5.52)
tenure	0.036*** (4.49)	0.032*** (4.04)	0.033*** (4.10)
rns	-0.079*** (-2.62)	-0.085*** (-2.87)	-0.080*** (-2.66)
smsa	0.138*** (4.92)	0.132*** (4.65)	0.131*** (4.61)
_cons	4.058*** (44.13)	4.039*** (42.70)	4.002*** (40.98)
N	758	758	758
adj. R-sq	0.349	0.351	0.352

t statistics in parentheses

* p<0.1, ** p<0.05, *** p<0.01

- ⑧ Check instrument validity
 - Run and examine the **reduced form**.
 - ▶ Reduced form is proportional to the casual effect of interest and is unbiased(OLS), so we should see the causal relation in the reduced form.
 - ▶ If you can't see the causal relation in the reduced form, it's probably not there.
 - ▶ **Notice!** Probably **med** is not exogenous enough (or not a very good IV).

- 8 Check instrument validity
- Check **exclusion restriction** : Overidentification test.
 - ▶ Instrumental variables (**med kww**) satisfy exogeneity.

```
. qui ivregress 2sls lw s expr tenure rns smsa (iq=med kww), r
. estat overid
Test of overidentifying restrictions:
Score chi2(1)          =  .151451  (p = 0.6972)
```

IV estimator in Stata

- 9 Check 2SLS estimates with LIML.
- LIML is less than precise than 2SIS but also less biased.
- If the results come out similar, be happy.

```
. ivregress liml lw s expr tenure rns smsa (iq=med kww), r
Instrumental variables (LIML) regression      Number of obs   =       758
                                             Wald chi2(6)    =      369.62
                                             Prob > chi2     =       0.0000
                                             R-squared       =       0.2768
                                             Root MSE       =       .36454
```

lw	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
iq	.0139764	.0060681	2.30	0.021	.0020831	.0258697
s	.0606362	.019034	3.19	0.001	.0233303	.0979421
expr	.0433416	.0074185	5.84	0.000	.0288016	.0578816
tenure	.0296237	.008323	3.56	0.000	.0133109	.0459364
rns	-.0433875	.034529	-1.26	0.209	-.1110631	.0242881
smsa	.1271796	.0297599	4.27	0.000	.0688512	.185508
_cons	3.214994	.4001492	8.03	0.000	2.430716	3.999272

```
Instrumented:  iq
Instruments:  s expr tenure rns smsa med kww
```

IV estimator in Stata

- Put all estimates into one table.

```
qui reg lw s expr tenure rns smsa,r
est sto ols_noiq
qui reg lw iq s expr tenure rns smsa,r
est sto ols_iq
qui ivreg2 lw s expr tenure rns smsa (iq=med kww), r
est sto tsls
qui ivreg2 lw s expr tenure rns smsa (iq=med kww), r liml
est sto liml

esttab ols_noiq ols_iq tsls liml, mtitle ///
      star(* 0.1 ** 0.05 *** 0.01) ///
      b(%6.3f) nogap compress order(s iq) ///
      stats(rkf j jp N r2_a, labels("First-stage F-statistic" ///
      "Overidentifying restrictions J-test and P-value" N r2_a) layout(@ `"'@ (@)"'"' @ @) )
```


IV estimator in Stata

- 10 Put all estimates into one table.

	(1) ols_noiq	(2) ols_iq	(3) tsls	(4) liml
s	0.103*** (16.53)	0.093*** (13.30)	0.061*** (3.21)	0.061*** (3.19)
smsa	0.140*** (4.98)	0.137*** (4.92)	0.127*** (4.28)	0.127*** (4.27)
iq		0.003*** (2.90)	0.014** (2.31)	0.014** (2.30)
expr	0.038*** (5.76)	0.039*** (5.91)	0.043*** (5.85)	0.043*** (5.84)
tenure	0.036*** (4.45)	0.034*** (4.33)	0.030*** (3.56)	0.030*** (3.56)
rns	-0.084*** (-2.85)	-0.075** (-2.49)	-0.044 (-1.26)	-0.043 (-1.26)
_cons	4.104*** (46.81)	3.895*** (33.60)	3.218*** (8.08)	3.215*** (8.03)
First-st_c			13.403	13.403
Overiden_t			0.151 (0.697)	0.151 (0.697)
N	758.000	758.000	758.000	758.000
r2_a	0.348	0.355	0.272	0.271

t statistics in parentheses

* p<0.1, ** p<0.05, *** p<0.01

- Followed research

- ▶ Return on investment in education is the core issue of labor economics.
- ▶ Behrman et al(1980) compared **identical twins** with different years of education to control the factors such as genetics and family background.
- ▶ Angrist and Krueger(1991) used **the quarter of birth** as the instrumental variable of years of education.
- ▶ Bound et al(1995) found the quarter of birth is a **weak** instrument.
- ▶ Buckles and Hungerman(2012)'s latest research showed that the quarter of birth is **not** independent of family background.

Section 2

RDD in Stata

Subsection 1

Review the Theory

- Review the Theory : Summary

In a Summary

RDD in the toolkit of Causal Inference

- It is so called the **nearest** method to RCT which identify causal effect of treatment on outcome.
- RDD needs a arbitrary cut-off and agents can **imperfect** manipulate the treatment.
- Two types
 - Sharp RD
 - Fuzzy RD
- Assumption: continued at the cut-off
- Concerns:
 - Functional form
 - Bandwidth selection
 - Bin selection

- Review the Theory : Main idea

Main Idea of Regression Discontinuity Design

- Regression Discontinuity Design (RDD) exploits the facts that:
 - Some rules are *arbitrary* and generate a *discontinuity* in treatment assignment.
 - The treatment assignment is determined based on whether a unit exceeds some threshold on a variable (**assignment variable**, **running variable** or **forcing variable**)
 - Assume other factors *do NOT change* abruptly at threshold.
 - Then any change in outcome of interest can be attributed to the assigned treatment.

- Review the Theory : Two types

RDD: Theory and Application

Sharp RDD and Fuzzy RDD

- In general, depending on enforcement of treatment assignment, RDD can be categorized into two types:
 - ① **Sharp RDD**: nobody below the cutoff gets the “treatment”, everybody above the cutoff gets it
 - Everyone follows treatment assignment rule (all are compliers).
 - Local randomized experiment with perfect compliance around cutoff.
 - ② **Fuzzy RDD**: the probability of getting the treatment jumps discontinuously at the cutoff (NOT jump from 0 to 1)
 - Not everyone follows treatment assignment rule.
 - Local randomized experiment with partial compliance around cutoff.
 - Using initial assignment as an instrument for actual treatment.

- Review the Theory : Assumption

- **Deterministic Assumption**

$$D_i = 1(X_i \geq c)$$

- Treatment assignment is a deterministic function of the assignment variable X_i and the threshold c .

- **Continuity Assumption**

- $E[Y_{1i}|X_i]$ and $E[Y_{0i}|X_i]$ are continuous at $X_i = c$
 - Assume potential outcomes do not change at cutoff.
 - This means that except treatment assignment, all other unobserved determinants of Y_i are continuous at cutoff c .
 - This implies no other confounding factor affects outcomes at cutoff c .
 - Any observed discontinuity in the outcome can be attributed to treatment assignment.

- Review the Theory : Identification in Sharp RD

Sharp RDD specification

- A simple RD regression is

$$Y_i = \alpha + \rho D_i + \gamma(X_i - c) + u_i$$

- Y_i is the outcome variable
- D_i is the the treatment variable(indepent variable)
- X_i is the running variable
- c is the value of cut-off
- u_i is the error term including other factors
- **Question:** Which parameter do we care about the most?

- Review the Theory : Identification in Sharp RD

- More generally, we could also estimate two separate regressions for each side respectively.

$$Y_i^b = \beta^b + f(X_i^b - c) + u_i^b$$

$$Y_i^a = \beta^a + g(X_i^a - c) + u_i^a$$

- Can do all in one step; just use all the data at once and estimate:

$$Y_i = \alpha + \rho D_i + f(X_i - c) + D_i \times h(X_i - c) + u_i$$

where D_i is a dummy variable for treated status.

- Review the Theory : Identification in Fuzzy RD

- Encourage Variable:

$Z_i = 1$ if assign to treatment group

$Z_i = 0$ if assign to control group

- Then the **First Stage** of FRD regression:

$$P(D_i = 1|x_i) = \alpha_1 + \phi Z_i + f(x_i - c) + Z_i \times g(x_i - c) + \eta_{1i}$$

- The **second stage** regression is

$$Y_i = \alpha_2 + \delta \hat{D}_i + f(x_i - c) + \hat{D}_i \times g(x_i - c) + \eta_{2i}$$

- The **reduced form** regression in FRD is

$$Y_i = \alpha_3 + \beta Z_i + f(x_i - c) + Z_i \times g(x_i - c) + \eta_{3i}$$

Subsection 2

Introduction : Package & Commands

- Introduction : Package & Commands

- ▶ Package : [Install Link](#)
- ▶ **rdrobust** package : inference and graphical procedures using local polynomial and partitioning regression methods.
 - rdrobust*- : Local Polynomial Regression Discontinuity Estimation with Robust Bias-Corrected Confidence Intervals and Inference Procedures.
 - rdbwselect*- : Data-driven Bandwidth Selection,
 - rdplot*- : Data-Driven Regression Discontinuity Plots.
- ▶ **rddensity** package : manipulation testing using local polynomial density methods.
- ▶ **Others** : -*cmogram*- ; -*rd*- ; -*rdcv*- ; -*DCdensity*- ; ...

Subsection 3

Example for Sharp RDD

- Example for Sharp RDD

- ▶ Data : the dataset comes from a study on **party advantages** in **U.S. Senate elections** for the period 1914–2010.
- ▶ We focus here on the RD effect of **the Democratic party winning a U.S. Senate seat** on **the vote share obtained in the following election for that same seat**.
- ▶ The unit of observation is the **state**.
- ▶ Main variables :
 - demmv* : ranges from -100 to 100 and records the Democratic party's margin of victory in the statewide election for a given U.S. Senate seat (the vote share of the Democratic party — the vote share of its strongest opponent).
 - demvoteshfor2* : ranges from 0 to 100 and records the Democratic vote share in the following election for the same seat.
- ▶ To estimate the **incumbency advantage of parties** with an RD design.

RDD in Stata

- Example for Sharp RDD
 - ▶ Re-labeling the three main variables

```
. use senate, clear
```

```
. sum
```

Variable	Obs	Mean	Std. Dev.	Min	Max
state	1,390	40.01367	21.99304	1	82
year	1,390	1964.63	28.05466	1914	2010
dopen	1,380	.2471014	.4314826	0	1
population	1,390	3827919	4436950	78000	3.73e+07
presdemvot_1	1,387	46.11975	14.31701	0	97.03408
demmv	1,390	7.171159	34.32488	-100	100
demvoteshl_1	1,349	52.69048	18.2706	0	100
demvoteshl_2	1,308	52.86918	18.23913	0	100
demvoteshf_1	1,341	52.41856	18.36641	0	100
demvoteshf_2	1,297	52.66627	18.12219	0	100
demwinprv1	1,349	.5441067	.4982355	0	1
demwinprv2	1,308	.543578	.4982879	0	1
dmidterm	1,390	.5136691	.499993	0	1
dpresdem	1,390	.3884892	.4875822	0	1

- Example for Sharp RDD

- ▶ Re-labeling the three main variables

- Assignment variable (running variable) : X

- Outcome variable: Y

- Treatment variable : T

- Threshold (cutoff) for treatment assignment : $c=0$

- Example for Sharp RDD

- ▶ Re-labeling the three main variables

```
. rename demmv X           //X--民主党获胜的差额
. rename demvoteshfor2 Y   //Y--t+2期民主党得票数

. gen T=.
. replace T=0 if X<0 & X!=.
. replace T=1 if X>=0 & X!=.
/* the Democratic party wins the election for that seat.*/

. label var T "Democratic Win at t"
```

- Example for Sharp RDD
 - ▶ Check RD's type

```
. gen ranwin=(X>=0)  
. tab ranwin T
```

ranwin	Democratic Win at t		Total
	0	1	
0	640	0	640
1	0	750	750
Total	640	750	1,390

- Example for Sharp RDD

- ▶ Show the scatter plot of the raw data (where each point is an observation).

```
. twoway (scatter Y X, msize(vsmall)          ///  
         mcolor(black) xline(0, lcolor(black))),  ///  
         graphregion(color(white)) ytitle(Outcome)  ///  
         xtitle(Score)  
  
. graph export fig1.png, width(500) replace  
  (file fig1.png written in PNG format)
```

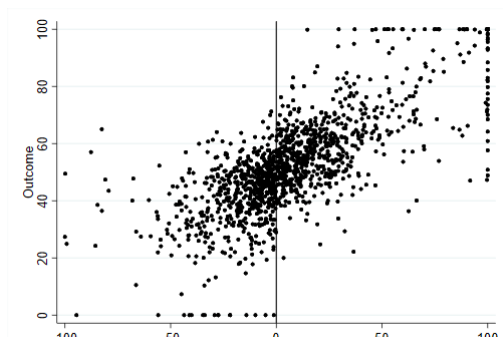
RDD in Stata

- Example for Sharp RDD

- ▶ Show the scatter plot of the raw data (where each point is an observation).
- ▶ Often hard to see "jumps" or discontinuities in the outcome-score relationship by simply looking at the raw data
- ▶ Two problems :

样本太多时不够直观；

实际分析时中跳跃现象可能不那么清晰。



- Example for Sharp RDD
 - ▶ Three Steps:
 - ① Graph the data for visual inspection
 - ② Estimate the treatment effect using regression methods
 - ③ Run checks on assumptions underlying research design

Subsection 4

Example for Sharp RDD : Step 1

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ A more useful approach is to aggregate or “smooth” the data before plotting.
 - ▶ The typical RD plot presents two ingredients :
 - (i) a global polynomial fit, represented by a **solid line**, using the original **raw** data.
 - (ii) local sample means, represented by **dots**, choosing bins of the score, calculating the mean of the outcome for the observations falling within each bin, and then plotting the **average outcome in each bin** against the mid point of the bin.

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*

```
preserve
rdplot Y X, nbins(20 20) genvars support(-100 100)
gen obs = 1
collapse (mean) rdplot_mean_x rdplot_mean_y (sum) obs, by (rdplot_id)
order rdplot_id
tabstat rdplot_mean_x rdplot_mean_y obs,by(rdplot_id)
restore
```

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (1) : Choosing the Location of Bins
 - ① **Evenly-spaced bins** : bins that have equal length.

```
. rdplot Y X, nbins(20 20) binsselect(es)    ///  
    graph_options(graphregion(color(white))  ///  
    xtitle(Score) ytitle(Outcome))
```

RD Plot with RD plot with manually set number of bins.

Cutoff c = 0	Left of c	Right of c	Number of obs =	1297
Number of obs	595	702	Kernel =	Uniform
Eff. Number of obs	595	702		
Order poly. fit (p)	4	4		
BW poly. fit (h)	100.000	100.000		
Number of bins scale	1.000	1.000		

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (1) : Choosing the Location of Bins
 - ① **Evenly-spaced bins** : bins that have equal length.

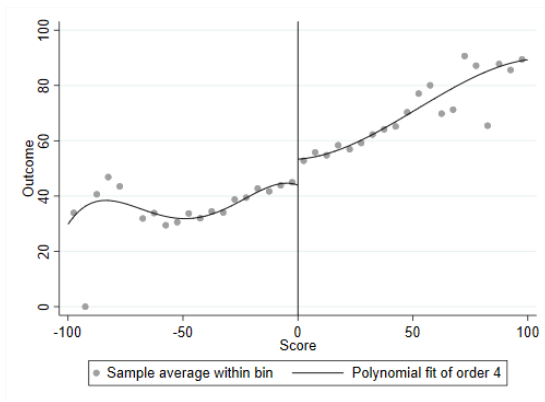
```
Outcome: Y. Running variable: X.
```

	Left of c	Right of c
Bins selected	20	20
Average bin length	5.000	5.000
Median bin length	5.000	5.000
IMSE-optimal bins	8	9
Mimicking Var. bins	15	35
Rel. to IMSE-optimal:		
Implied scale	2.500	2.222
WIMSE var. weight	0.060	0.084
WIMSE bias weight	0.940	0.916

```
. graph export fig2.png,width(500) replace  
(note: file fig2.png not found)  
(file fig2.png written in PNG format)
```

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (1) : Choosing the Location of Bins
 - ① **Evenly-spaced bins** : bins that have equal length.



Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (1) : Choosing the Location of Bins
 - ② **Quantile-spaced bins** : bins that contain (roughly) the same number of observations.

```
. rdplot Y X, nbins(20 20) binselect(qs) ///  
  graph_options(graphregion(color(white))) ///  
  xtitle(Score) ytitle(Outcome)
```

RD Plot with RD plot with manually set number of bins.

Cutoff $c = 0$	Left of c	Right of c	Number of obs =	1297
Number of obs	595	702	Kernel =	Uniform
Eff. Number of obs	595	702		
Order poly. fit (p)	4	4		
BW poly. fit (h)	100.000	100.000		
Number of bins scale	1.000	1.000		

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*

- ▶ Bin selection (1) : Choosing the Location of Bins

- ② **Quantile-spaced bins** : bins that contain (roughly) the same number of observations.

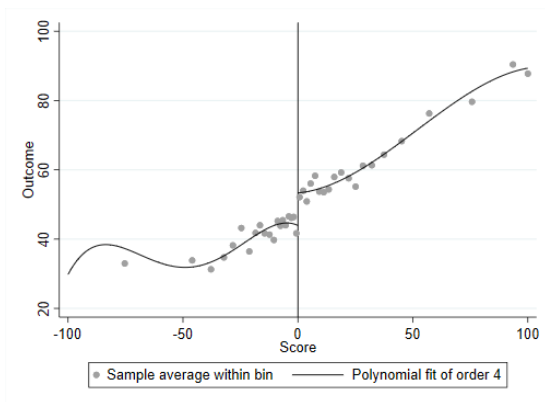
```
Outcome: Y. Running variable: X.
```

	Left of c	Right of c
Bins selected	20	20
Average bin length	5.000	5.000
Median bin length	1.912	2.771
IMSE-optimal bins	21	16
Mimicking Var. bins	28	49
Rel. to IMSE-optimal:		
Implied scale	0.952	1.250
WIMSE var. weight	0.537	0.339
WIMSE bias weight	0.463	0.661

```
. graph export fig3.png,width(500) replace  
(note: file fig3.png not found)  
(file fig3.png written in PNG format)
```

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (1) : Choosing the Location of Bins
 - ② **Quantile-spaced bins** : bins that contain (roughly) the same number of observations.



Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ① Integrated Mean Squared Error (IMSE) Method
 - If we choose a large number of bins (narrower) :
 - small bias** – the bins are smaller and the **local constant** fit is better.
 - less precisely** – less observations **per bin**, thus more variability within bin.
 - balance squared-bias and variance so that the IMSE is (approximately) minimized.

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ① Integrated Mean Squared Error (IMSE) Method

```
. rdplot Y X, binselect(es)          ///
      graph_options(graphregion(color(white)))  ///
      xtitle(Score) ytitle(Outcome)

/* The IMSE criterion leads to different numbers of ES bins above and
below the cutoff.*/
```

RD Plot with evenly spaced number of bins using spacings estimators.

Cutoff c = 0	Left of c	Right of c	Number of obs =	1297
			Kernel =	Uniform
Number of obs	595	702		
Eff. Number of obs	595	702		
Order poly. fit (p)	4	4		
BW poly. fit (h)	100.000	100.000		
Number of bins scale	1.000	1.000		

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ① Integrated Mean Squared Error (IMSE) Method

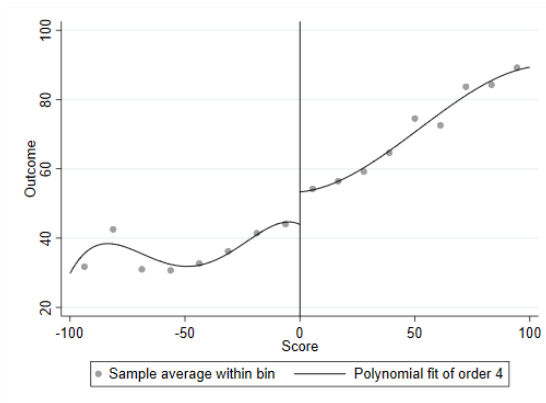
Outcome: Y. Running variable: X.

	Left of c	Right of c
Bins selected	8	9
Average bin length	12.500	11.111
Median bin length	12.500	11.111
IMSE-optimal bins	8	9
Mimicking Var. bins	15	35
Rel. to IMSE-optimal:		
Implied scale	1.000	1.000
WIMSE var. weight	0.500	0.500
WIMSE bias weight	0.500	0.500

```
. graph export fig4.png,width(500) replace
(note: file fig4.png not found)
(file fig4.png written in PNG format)
```

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ① Integrated Mean Squared Error (IMSE) Method



Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ① Integrated Mean Squared Error (IMSE) Method

```
rdplot Y X, binselect(qs)           ///  
      graph_options(graphregion(color(white)))  ///  
      xtitle(Score) ytitle(Outcome))
```

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ② Mimicking Variance (MV) Method
 - “mimics” the overall variability in the raw scatter plot of the data.
 - MV method leads to a larger number of bins than the IMSE method.
 - More dots representing local means, thus giving a better sense of the variability of the data.

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ② Mimicking Variance (MV) Method

```
. rdplot Y X, binselect(esmv)          ///Default
  graph_options(graphregion(color(white)) ///
  xtitle(Score) ytitle(Outcome))
```

RD Plot with evenly spaced mimicking variance number of bins using spacings estimated

Cutoff $c = 0$	Left of c	Right of c	Number of obs =	1297
Number of obs	595	702	Kernel =	Uniform
Eff. Number of obs	595	702		
Order poly. fit (p)	4	4		
BW poly. fit (h)	100.000	100.000		
Number of bins scale	1.000	1.000		

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ② Mimicking Variance (MV) Method

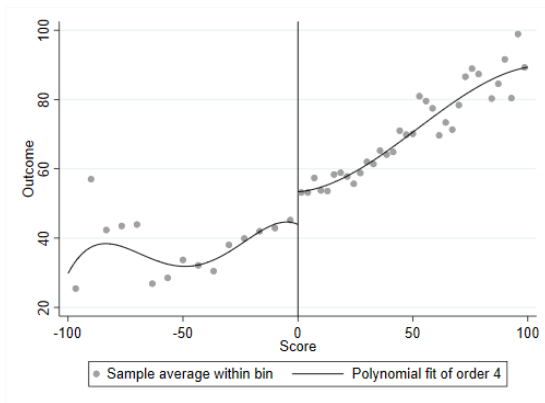
Outcome: Y. Running variable: X.

	Left of c	Right of c
Bins selected	15	35
Average bin length	6.667	2.857
Median bin length	6.667	2.857
IMSE-optimal bins	8	9
Mimicking Var. bins	15	35
Rel. to IMSE-optimal:		
Implied scale	1.875	3.889
WIMSE var. weight	0.132	0.017
WIMSE bias weight	0.868	0.983

```
. graph export fig5.png,width(500) replace  
(file fig5.png written in PNG format)
```

Example for Sharp RDD

- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ② Mimicking Variance (MV) Method



Example for Sharp RDD

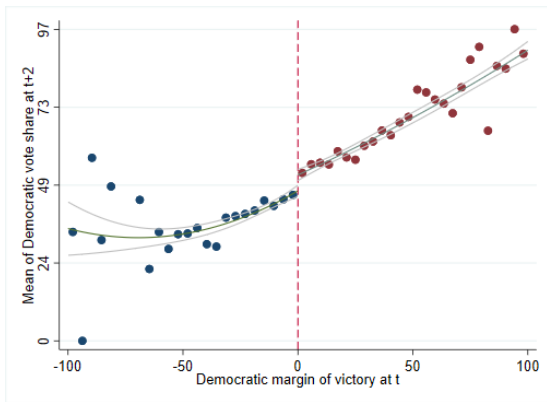
- RDD graphical analysis : *-rdplot-*
 - ▶ Bin selection (2) : Choosing the Number of Bins
 - ② Mimicking Variance (MV) Method

```
rdplot Y X, binselect(qsmv)          ///  
    graph_options(graphregion(color(white)) ///  
    xtitle(Score) ytitle(Outcome))
```

Example for Sharp RDD

- RDD graphical analysis : `-cmogram-`

```
. cmogram Y X, cut(0) scatter lineat(0) qfitci  
  
. graph export fig6.png,width(500) replace  
(note: file fig6.png not found)  
(file fig6.png written in PNG format)
```



Subsection 5

Example for Sharp RDD : Step 2

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - There are 2 types of strategies for correctly specifying the functional form in a RDD:
 - 1 **Parametric/global method:** Use all available observations and estimate treatment effects based on a specific functional form for the outcome and assignment variable relationship.
 - 2 **Nonparametric/local method:** Use the observations around cutoff: Compare the outcome of treated and untreated observations that lie within specific bandwidth.

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Parametric/Global Approach (全局多项式回归)

```
sum X
local hvalueR=r(max)
local hvalueL= abs(r(min))

rdrobust Y X, h(`hvalueL' `hvalueR') //自动选择阶数
rdrobust Y X, h(`hvalueL' `hvalueR') p(2) //二阶拟合
rdrobust Y X, h(`hvalueL' `hvalueR') p(3) //三阶拟合
```

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)
 - ▶ 三种方法 (任选):
 - 方法一: standard least-squares estimation (OLS)
 - 方法二: -rdrobust-进行的非参数估计
 - 方法三: -rd-进行的非参数估计

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - Nonparametric/Local Approach : local linear regression (局部线性回归)
 - 方法一: standard least-squares estimation (OLS)

```
. rdbwselect Y X, c(0) kernel(uni) bwselect(mserd) //选择最优带宽h
Bandwidth estimators for sharp RD local polynomial regression.
```

Cutoff c =	Left of c	Right of c	Number of obs =	1297
Number of obs	595	702	Kernel =	Uniform
Min of X	-100.000	0.036	VCE method =	NN
Max of X	-0.079	100.000		
Order est. (p)	1	1		
Order bias (q)	2	2		

Outcome: Y. Running variable: X.

Method	BW est. (h)		BW bias (b)	
	Left of c	Right of c	Left of c	Right of c
mserd	11.597	11.597	22.944	22.944

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法一: standard least-squares estimation (OLS)

```
preserve
keep if X>=-11.597 & X<=11.597

local i=1
forvalues i=2/4
  gen X`i`=X^`i'                                     // 产生分配变量的平方、三次方、四次方

eststo x1 : qui reg Y 1.T, r
eststo x2 : qui reg Y T##c.X, r                       //局部线性回归法, 一阶
eststo x3 : qui reg Y T##c.(X X2), r                 //局部线性回归法, 选择2阶多项式
eststo x4 : qui reg Y T##c.(X X2 X3), r             //局部线性回归法, 选择3阶多项式
eststo x5 : qui reg Y T##c.(X X2 X3 X4), r          //局部线性回归法, 选择4阶多项式

esttab x1 x2 x3 x4 x5,                               ///
      star(* .1 ** .05 * .01)                       ///
      nogap nonumber replace                         ///
      drop(0.T*) se(%5.4f) ar2 aic(%10.4f) bic(%10.4f)
restore
```


Example for Sharp RDD

- Estimate the treatment effect using regression methods

- ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法一: standard least-squares estimation (OLS)

	Y	Y	Y	Y	Y
1.T	9.762* (0.8349)	7.202* (1.6332)	8.832* (2.3939)	13.27* (3.1258)	16.08* (3.8665)
X		0.240 (0.2028)	-1.670** (0.7598)	-2.793 (1.8143)	-7.721** (3.7073)
1.T#c.X		-0.0122 (0.2585)	3.029* (0.9956)	0.566 (2.3214)	6.207 (4.7669)
X2			-0.169** (0.0667)	-0.405 (0.3493)	-2.285* (1.3548)
1.T#c.X2			0.0691 (0.0860)	1.110** (0.4670)	2.708 (1.7333)
X3				-0.0135 (0.0196)	-0.265 (0.1819)
1.T#c.X3				-0.0341 (0.0267)	0.256 (0.2288)
X4					-0.0109 (0.0079)
1.T#c.X4					0.00921 (0.0099)
_cons	44.28* (0.6010)	45.60* (1.2794)	41.91* (1.8869)	40.71* (2.5757)	37.51* (3.1696)
N	506	506	506	506	506
adj. R-sq	0.209	0.211	0.224	0.230	0.231
AIC	3709.6679	3710.3097	3704.0478	3702.2052	3703.7348
BIC	3718.1210	3727.2159	3729.4070	3736.0175	3746.0002

Standard errors in parentheses

* p<.1, ** p<.05, * p<.01

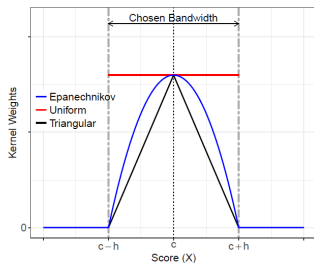
Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法二: *-rdrobust*-进行的非参数估计

p : set the order of the polynomial. Default is $p(1)$.

kernel : set the kernel. Default is `kernel(triangular)`.



h : choose the bandwidth manually.

c : sets the RD cutoff. Default is $c(0)$.

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法二: -*rdrobust*-进行的非参数估计

`bwselect()` : bandwidth selection procedure to be used. Default is `bwselect(mserd)`.

If a smaller h :

fewer observations—increase the variance of the estimated coefficients.

local polynomial approximation—will reduce treatment effect bias.

MSE : bias-variance trade-off.

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法二: -*rdrobust*-进行的非参数估计

```
rdrobust Y X, kernel(uniform) p(1)
rdrobust Y X, c(0) kernel(uni) bwselect(mserd) p(2) h(11.597) all
rdrobust Y X, c(0) kernel(uni) bwselect(mserd) p(3) h(11.597) all
rdrobust Y X, c(0) kernel(uni) bwselect(mserd) p(4) h(11.597) all
```

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

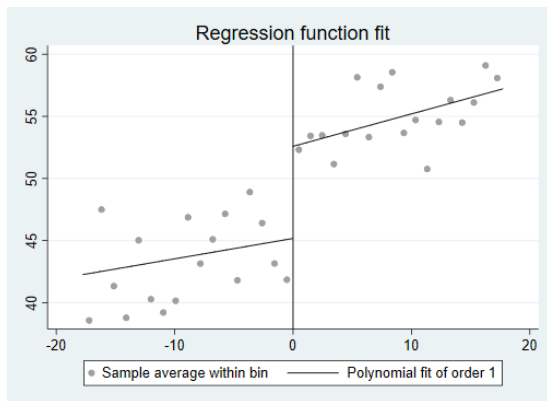
方法二: *-rdrobust-*进行的非参数估计

```
* Using rdrobust and showing the associated rdplot
. rdrobust Y X, p(1) kernel(triangular) bwselect(mserd)
. eret list
. local bandwidth = e(h_1)
. rdplot Y X if abs(X) <= `bandwidth`, p(1) h(`bandwidth`) kernel(triangular)
. graph export fig7.png,width(500) replace
```

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法二: *-rdrobust*-进行的非参数估计



Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法三: *-rd-*进行的非参数估计

```
. rd Y X, mbw(100) gr z0(0) kernel(tri) //给出了带宽取最优带宽50%和200%的回归结果
```

```
Two variables specified; treatment is assumed to jump from zero to one at Z=0.
```

```
Assignment variable Z is X
```

```
Treatment variable X_T unspecified
```

```
Outcome variable y is Y
```

```
Command used for lpoly; Kernel used: triangle (default)
```

```
Bandwidth: 7.5496767; loc Wald Estimate: 9.6449759
```

```
(93 missing values generated)
```

```
(93 missing values generated)
```

```
(93 missing values generated)
```

```
Estimating for bandwidth 7.549676665805968
```

Y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
lwald	9.644976	2.1155	4.56	0.000	5.498673	13.79128

```
. graph export fig8.png,width(500) replace
```

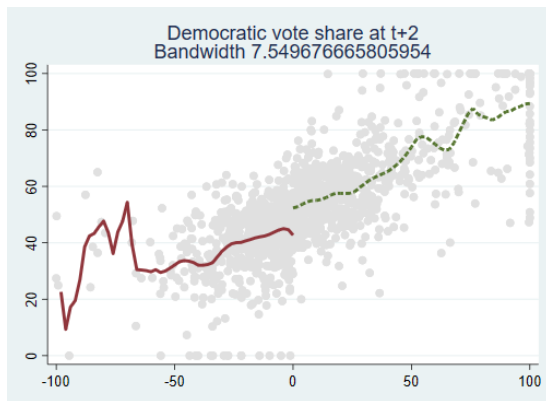
```
(note: file fig8.png not found)
```

```
(file fig8.png written in PNG format)
```

Example for Sharp RDD

- Estimate the treatment effect using regression methods
 - ▶ Nonparametric/Local Approach : local linear regression (局部线性回归)

方法三: -rd-进行的非参数估计



Subsection 6

Example for Sharp RDD : Step 3

- Testing the Validity of the RDD
 - ① **Test involving covariates(Nonoutcome Variable)** : Test whether other covariates exhibit a jump at the discontinuity
 - ② **Test sorting behavior** : Testing discontinuity in the density of assignment variable X
 - ③ **Falsification Tests** :
 - ① Placebo Cutoffs
 - ② Sensitivity to Observations near the Cutoffs
 - ③ Sensitivity to Bandwidth Choice

Example for Sharp RDD

- Testing the Validity of the RDD

- 1 Test involving covariates(Nonoutcome Variable)

Using rdbwselect with covariates.

```
. global covariates "presdemvoteshlag1 demvoteshlag1 demvoteshlag2 demwinprv1 demwinprv2 dmidte rm dpresdem"
. rdrobust Y X, covs($covariates) p(1) kernel(tri) bwselect(mserd)
```

Covariate-adjusted sharp RD estimates using local polynomial regression.

Cutoff c = 0	Left of c	Right of c	Number of obs =	1213
Number of obs	555	658	BW type	= mserd
Eff. Number of obs	326	295	Kernel	= Triangular
Order est. (p)	1	1	VCE method	= NN
Order bias (q)	2	2		
BW est. (h)	17.266	17.266		
BW bias (b)	27.178	27.178		
rho (h/b)	0.635	0.635		

Outcome: Y. Running variable: X.

Method	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
Conventional	7.0876	1.4767	4.7995	0.000	4.19327 9.98194
Robust	-	-	4.0449	0.000	3.67067 10.572

Covariate-adjusted estimates. Additional covariates included: 7

Example for Sharp RDD

- Testing the Validity of the RDD

- ① **Test involving covariates(Nonoutcome Variable)**

Test whether other covariates exhibit a jump at the discontinuity.

- * There should be no jump in other covariates.
- * 从图形，似乎是存在跳跃的，但这并不严格，要看回归结果

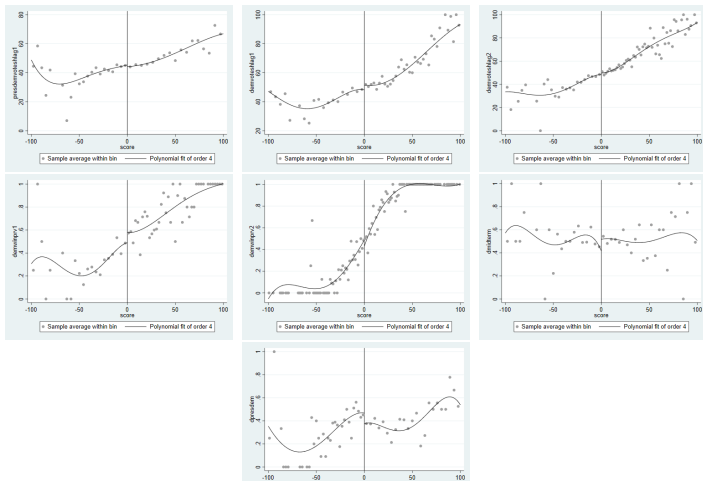
```
^^Iforeach y of global covariates {  
^^I^^Iqui rdplot `y' X, graph_options(xtitle("score")) saving(`y')  
^^I  
^^I^^Igraph export fig_`y'.png, width(500) replace  
^^I}
```

Example for Sharp RDD

- Testing the Validity of the RDD

- 1 Test involving covariates(Nonoutcome Variable)

Test whether other covariates exhibit a jump at the discontinuity.



Example for Sharp RDD

- Testing the Validity of the RDD

- ① **Test involving covariates(Nonoutcome Variable)**

Test whether other covariates exhibit a jump at the discontinuity.

* 估计具体系数看是否显著

```
. est clear
^^Iforeach y of global covariates {
^^I^^Ieststo : qui rdrobust `y' X, all
^^I}

. estab est1 est2 est3 est4 est5 est6 est7 , ///
      se r2 mtitle star(* 0.1 ** 0.05 *** 0.01) compress
```

Example for Sharp RDD

- Testing the Validity of the RDD

- 1 Test involving covariates(Nonoutcome Variable)**

Test whether other covariates exhibit a jump at the discontinuity.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
	est1	est2	est3	est4	est5	est6	est7
Conventi_1	-1.363 (1.383)	2.459 (2.052)	1.001 (1.918)	0.0728 (0.0724)	-0.0270 (0.0712)	0.0696 (0.0662)	-0.1000 (0.0714)
Bias-cor_d	-1.193 (1.383)	2.898 (2.052)	1.495 (1.918)	0.0773 (0.0724)	-0.0386 (0.0712)	0.0828 (0.0662)	-0.102 (0.0714)
Robust	-1.193 (1.633)	2.898 (2.454)	1.495 (2.246)	0.0773 (0.0866)	-0.0386 (0.0845)	0.0828 (0.0772)	-0.102 (0.0854)
N	1387	1349	1308	1349	1308	1390	1390
R-sq							

Standard errors in parentheses

* p<0.1, ** p<0.05, *** p<0.01

Example for Sharp RDD

- Testing the Validity of the RDD

- ② **Test sorting behavior**

Testing discontinuity in the density of assignment variable X
-*rddensity*-

```
. rdrobust Y X  
. local h = e(h_1) //获取最优带宽  
. rddensity X, p(1) h(`h' `h') plot
```

RD Manipulation Test using local polynomial density estimation.

Cutoff c = 0	Left of c	Right of c		
Number of obs	640	750	Number of obs =	1390
Eff. Number of obs	377	346	Model =	unrestricted
Order est. (p)	1	1	BW method =	manual
Order bias (q)	2	2	Kernel =	triangular
BW est. (h)	17.754	17.754	VCE method =	jackknife

Running variable: X.

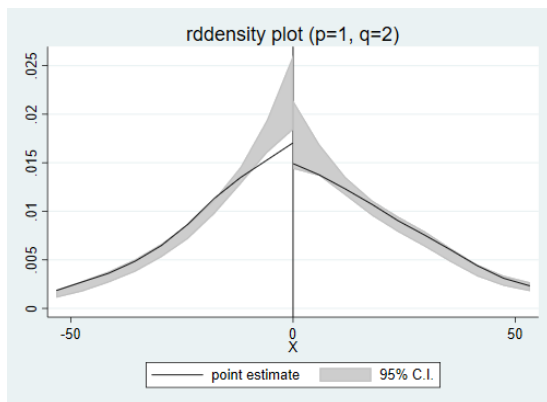
Method	T	P> T
Robust	-1.5083	0.1315

Example for Sharp RDD

- Testing the Validity of the RDD

② Test sorting behavior

Testing discontinuity in the density of assignment variable X
-rddensity-



Example for Sharp RDD

- Testing the Validity of the RDD

- ② **Test sorting behavior**

Testing discontinuity in the density of assignment variable X

Histogram(直方图)

```
. qui rddensity X
. local bandwidth_left = e(h_l)
. local bandwidth_right = e(h_r)

. twoway (histogram X if X >= `bandwidth_left' & X < 0, freq width(1) color(blue) lcolor(black) lwidth(vthin)) ///
        (histogram X if X >= 0 & X <= `bandwidth_right', freq width(1) color(red) lcolor(black) lwidth(vthin)), ///
        xlabel(-20(10)30) graphregion(color(white)) xtitle(Score) ytitle(Number of Observations) legend(off)

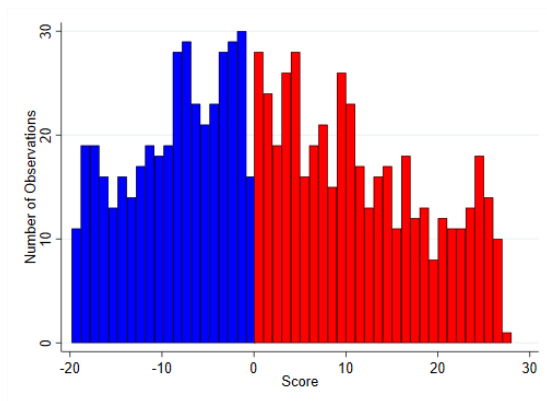
. graph export fig10.png, width(500) replace
(file fig10.png written in PNG format)
```

Example for Sharp RDD

- Testing the Validity of the RDD

- ② **Test sorting behavior**

Testing discontinuity in the density of assignment variable X
Histogram(直方图)



Example for Sharp RDD

- Testing the Validity of the RDD

- ② Test sorting behavior

Testing discontinuity in the density of assignment variable X
a more formal test : **McCrary(2008) test** -*DCdensity*-

```
. preserve
. DCdensity X, breakpoint(0) generate(Xj Yj r0 fhat se_fhat) // McCrary test

Using default bin size calculation, bin size = 1.84133021
Using default bandwidth calculation, bandwidth = 25.8493835
Discontinuity estimate (log difference in height): -.100745626
                                                    (.117145041)

Performing LLR smoothing.
110 iterations will be performed
.....
. gen t= -.100745626/.117145041 // 产生t值, 这个需要你根据系数提取出来
. display 2*ttail(50, t) // 得到p值, 50是自由度
1.6061102

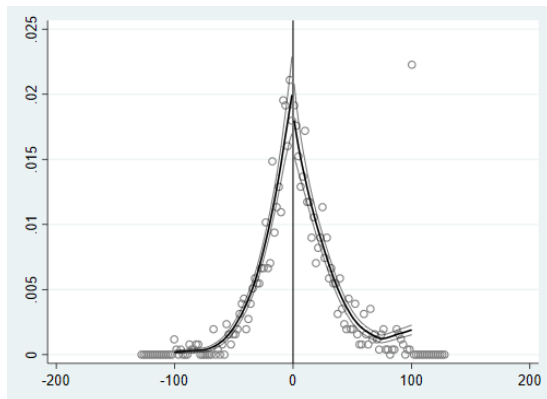
. graph export fig11.png, width(500) replace
(note: file fig11.png not found)
(file fig11.png written in PNG format)
. restore
. /**可以看出在5%显著性水平下实际上McCrary检验是通不过的, 证明没有操纵**/
```

Example for Sharp RDD

- Testing the Validity of the RDD

- ② **Test sorting behavior**

Testing discontinuity in the density of assignment variable X
a more formal test : **McCrary(2008) test** - *DCdensity*-



Example for Sharp RDD

- Testing the Validity of the RDD

- 3 Falsification Tests**

- ▶ Check 1 : Placebo Cutoffs

- ▶ 选择一个不同于断点的值作为安慰剂断点 (placebo cutoff points), 分别取真实断点两侧 25%、50%、75% 样本分位数处作为断点。

```
. sum X
+-----+-----+-----+-----+-----+
| Variable | Obs | Mean | Std. Dev. | Min | Max |
+-----+-----+-----+-----+-----+
| X        | 1,390 | 7.171159 | 34.32488 | -100 | 100 |
+-----+-----+-----+-----+-----+
. local xmax=r(max)
. local xmin=r(min)

**for values i=1(1)3{
**I**local jr=`xmax'/(4/(4-`i`))
**I**local jl=`xmin'/(4/(4-`i`))
**I**Iqui rdrobust Y X if X>0, c(`jr`)
**I**Iest store jl`i`
**I**Iqui rdrobust Y X if X<0, c(`jl`)
**I**Iest store jr`i`
**I}

. qui rdrobust Y X ,c(0) //加上真实断点的回归结果, 作为benchmark结果
. est store jbaseline
```

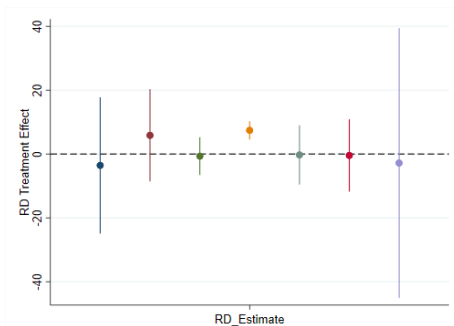
Example for Sharp RDD

- Testing the Validity of the RDD

- ③ **Falsification Tests**

- ▶ Check 1 : Placebo Cutoffs

```
. local vlist "j11 j12 j13 jbaseline jr3 jr2 jr1 "  
. coefplot `vlist', yline(0, lcolor(black) lpattern(dash)) drop(_cons) vertical ///  
    graphregion(color(white)) ytitle("RD Treatment Effect") legend(off)  
  
. graph export fig12.png, width(500) replace  
(file fig12.png written in PNG format)
```



Example for Sharp RDD

- Testing the Validity of the RDD

- ③ **Falsification Tests**

- ▶ Check 2 : Sensitivity to Observations near the Cutoffs

由于越接近断点的样本，越有动机去人为操控，删除最接近断点的样本，来观察回归是否显著（甜甜圈效应，donut hole approach）。

分别删除断点附近 1%，2%，3%，4% 和 5% 的样本，进行了 5 组稳健性检验。图形给出了回归系数和 95% 的置信区间。

```
. sum X
. local xmax=r(max)

^^Iforvalues i=1(1)5{
^^I^^Ilocal j=`xmax'*0.01*`i`
^^I^^Iqui rdrobust Y X if abs(X)>`j`
^^I^^Iest store obrob`i`
^^I}
```

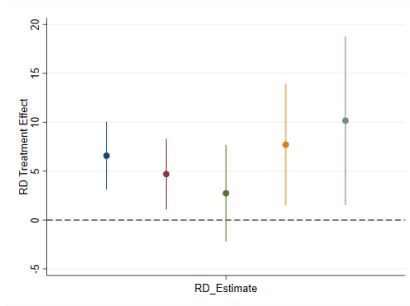

Example for Sharp RDD

- Testing the Validity of the RDD

- ③ Falsification Tests

- ▶ Check 2 : Sensitivity to Observations near the Cutoffs

```
. local vlist "obrob1 obrob2 obrob3 obrob4 obrob5"  
. coefplot `vlist', yline(0, lcolor(black) lpattern(dash)) drop(_cons) vertical ///  
    graphregion(color(white)) legend(off) ytitle("RD Treatment Effect")  
  
. graph export fig13.png, width(500) replace  
(note: file fig13.png not found)  
(file fig13.png written in PNG format)
```



Example for Sharp RDD

- Testing the Validity of the RDD

- ③ **Falsification Tests**

- ▶ Check 3 : Sensitivity to Bandwidth Choice

带宽长度会显著影响回归结果，一个稳健的结果要求对带宽长度不那么敏感。

提取最优带宽 h ，然后分别手动设置带宽为 h 的 25%-400% 倍，看回归结果是否仍旧显著。

图形给出了回归系数和 95% 的置信区间。

```
. qui rdrobust Y X //自动选择最优带宽
. local h = e(h_1) //获取最优带宽

^^Iforvalues i=1(1)8{
^^I^^Ilocal hrobust=`h'*0.25*`i'
^^I^^Iqui rdrobust Y X ,h(`hrobust`)
^^I^^Iest store hrob`i'
^^I}
```

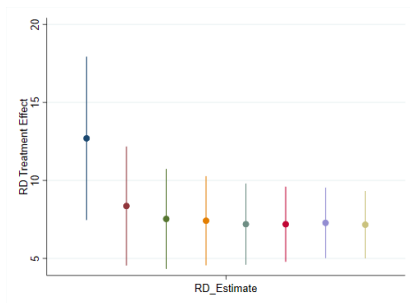
Example for Sharp RDD

- Testing the Validity of the RDD

- ③ Falsification Tests

- ▶ Check 3 : Sensitivity to Bandwidth Choice

```
. local vlist "hrob1 hrob2 hrob3 hrob4 hrob5 hrob6 hrob7 hrob8 "  
. coefplot `vlist', yline(0, lcolor(black) lpattern(dash)) drop(_cons) vertical ///  
  graphregion(color(white)) ytitle("RD Treatment Effect") legend(off)  
  
. graph export fig14.png, width(500) replace  
(note: file fig14.png not found)  
(file fig14.png written in PNG format)
```



Subsection 7

Example for Fuzzy RDD

Example for Fuzzy RDD

- 三种方法 (任选):
 - ▶ 方法一: `-rd-`
 - ▶ 方法二: `-rdrobust-`
 - ▶ 方法三: IV 估计

Example for Fuzzy RDD

- 方法一: `-rd-`

- ▶ Syntax

```
rd y d x, z0(real) strineq mbw(numlist) graph bdep oxline    ///  
    kernel(rectangle) covar(varlist) x(varlist)  
  
mbw(numlist)          //用来指定最优带宽的倍数, 默认值为mbw(50 100 200)  
z0(real)              //用来指定断点的位置, 默认值为z0(0), 即断点为原点  
*如果此处省去D, 则为SRD, 并根据分组变量X来计算处理变量  
graph                 //根据每一带宽, 画出局部线性回归图  
bdep                  //根据画图来考察断点回归估计量对带宽的依赖性  
oxline                //在此图的默认带宽上画出一条直线, 以便识别  
kernel(rectangle)    //使用均匀核 (uniform), 默认triangle  
covar(varlist)        //用来指定加入局部线性回归的协变量  
x(varlist)            //检验这些协变量在断点处是否存在跳跃 (估计跳跃值和
```

Example for Fuzzy RDD

- 方法一: *-rd-*

- ▶ Example background

在美国国会, 有一个民主党代表可能是被认为是对国会选区的一种 **treatment**。

美国国会选区, 如果有民主党众议员, 对该选区的联邦政府的开支具有一定影响。

传统意义上, 民主党会更倾向于政府, 如果当选, 会加大对联邦政府的开支。

然而直接对二者进行回归, 可能会遗漏变量问题或者双向因果关系。

为此选择该民主党候选人的得票比例作为分组变量, Z 是民主党候选人获得的选票份额。

以 0.5 为断点 (在民主党与共和党的政治中, $Z \geq 0.5$, 则当选, 反之落选), 进行 RDD。

Example for Fuzzy RDD

- 方法一: *-rd-*

- ▶ Example Data

```
. ssc inst rd, replace
. net get rd
. use votex, clear

* lne //选取联邦政府开支的对数
* d //分组变量, 民主党派候选人的得票比例减去0.5, 以标准化
* win //民主党派候选人当选
* 另外还包括一些协变量

. desc
```


Example for Fuzzy RDD

● 方法一: `-rd-`

▶ SRD

```
//OLS回归
reg lne win i votpop bla-vet
*回归结果虽然win表示当选了,会增加lne的支出,但是不显著

//选择默认的带宽以及triangle kernel进行RD
rd lne d, gr mbw(100)
*不显著,说明拥有民主党派候选人当选的选区并不能显著的增加联邦政府开支

//加入协变量进行RD,省略作图
rd lne d, mbw(100) cov(i votpop black blucllr farmer fedwrkr forborn manuf unemployd union urban veterans)
*显示估计值虽然为正,但是依然不显著

//去掉协变量,同时估计三种带宽,并画出估计值对带宽的依赖性
rd lne d, gr bdep oxline
*改变带宽对估计值有一定的影响,但是三个估计值全部为负,且依然不显著。可以看出,各个断点回归估计量对带宽的依赖性不大。

//检验协变量在断点处是否存在跳跃
rd lne d, mbw(100) x(i votpop black blucllr farmer fedwrkr forborn manuf unemployd union urban veterans)
*farmer的P值为 0.036,其余的协变量的条件密度函数在断点处都是连续的,即只有farmer(农民占人口比例)存在跳跃。
```

Example for Fuzzy RDD

- 方法一: `-rd-`

- ▶ FRD

```
. //生成一个新的处理变量randwin, 使得randwin不完全由分组变量d决定。  
. set seed 20181203  
. g byte randwin=cond(uniform(<.1,1-win,win)  
. tab randwin win
```

randwin	Dem Won Race		Total
	0	1	
0	123	14	137
1	8	204	212
Total	131	218	349

*结果显示randwin与win基本相同, 但不完全相同, 说明randwin不完全由分组变量d决定。

Example for Fuzzy RDD

- 方法一: `-rd-`

- ▶ FRD

```
. //使用最优带宽与默认的triangle kernel进行FRD(含协变量)
. rd lne randwin d, gr mbw(100) cov(i votpop black blucllr farmer fedwrkr forborn manuf unemply
> d union urban veterans)
```

Three variables specified; jump in treatment at Z=0 will be estimated. Local Wald Estimate is the ratio of jump in outcome to jump in treatment.

Assignment variable Z is d
Treatment variable X_T is randwin
Outcome variable y is lne

Command used for graph: lpoly; Kernel used: triangle (default)
Bandwidth: .29287776; loc Wald Estimate: -.09974965
Estimating for bandwidth .2928777592534943

A predicted value of treatment at cutoff lies outside feasible range;
switching to local mean smoothing for treatment discontinuity.

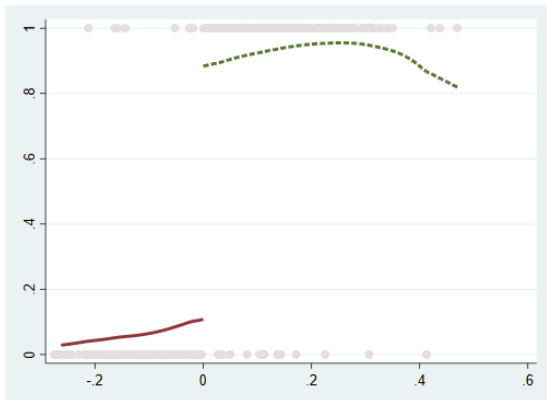
lne	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
numer	.0543733	.0900181	0.60	0.546	-.1220589	.2308055
denom	.8734363	.0301089	29.01	0.000	.814424	.9324487
lwald	.0622522	.1028807	0.61	0.545	-.1393903	.2638946

```
. graph export fig15.png, width(500) replace
(file fig15.png written in PNG format)
```

Example for Fuzzy RDD

● 方法一: `-rd-`

▶ FRD



Example for Fuzzy RDD

- 方法二: *-rdrobust-*

```
* options : fuzzy(fuzzyvar [sharpbw])
. rdrobust lne d, fuzzy(randwin)

Fuzzy RD estimates using local polynomial regression.
Cutoff c = 0 | Left of c | Right of c | Number of obs = 349
              |              |             | BW type = mserd
              |              |             | Kernel = Triangular
              |              |             | VCE method = NN
-----|-----|-----|-----|-----|-----|-----|
Number of obs |      131      |      218      |
Eff. Number of obs |      73      |      105      |
Order est. (p) |           1    |           1    |
Order bias (q) |           2    |           2    |
BW est. (h)    |      0.142    |      0.142    |
BW bias (b)    |      0.186    |      0.186    |
rho (h/b)     |      0.762    |      0.762    |
First-stage estimates. Outcome: randwin. Running variable: d.
-----|-----|-----|-----|-----|-----|
Method | Coef. | Std. Err. | z | P>|z| | [95% Conf. Interval]
-----|-----|-----|-----|-----|-----|
Conventional | .74178 | .09171 | 8.0888 | 0.000 | .562043 | .921522
Robust | - | - | 7.3050 | 0.000 | .53871 | .933789
Treatment effect estimates. Outcome: lne. Running variable: d. Treatment Status: randwin.
-----|-----|-----|-----|-----|-----|
Method | Coef. | Std. Err. | z | P>|z| | [95% Conf. Interval]
-----|-----|-----|-----|-----|-----|
Conventional | -.12768 | .20294 | -0.6292 | 0.529 | -.525433 | .270072
Robust | - | - | -0.6411 | 0.521 | -.665039 | .3372
```

Example for Fuzzy RDD

● 方法三: IV 估计

```
. use frd.dta, clear
. sum
```

Variable	Obs	Mean	Std. Dev.	Min	Max
anno	120	1998.667	3.831171	1993	2004
esse_m	120	0	6.230853	-10	10
mc	120	9.792913	.1016675	9.522602	9.995189
mcn	120	9.726134	.093931	9.46739	9.921432
mf	120	6.100074	.1004518	5.892969	6.368767
pen	120	.4036111	.3359836	0	.9861111
obsc	120	88.175	35.007	21	184
obscn	120	88.175	35.007	21	184
obsf	120	88.11667	34.99176	21	184
obsp	120	88.175	35.007	21	184
anno1993	120	-7.45e-09	.3742406	-.1666667	.8333333
anno1995	120	-7.45e-09	.3742406	-.1666667	.8333333
anno1998	120	-7.45e-09	.3742406	-.1666667	.8333333
anno2000	120	-7.45e-09	.3742406	-.1666667	.8333333
anno2002	120	-7.45e-09	.3742406	-.1666667	.8333333
anno2004	120	-7.45e-09	.3742406	-.1666667	.8333333
elig	120	.5	.5020964	0	1
esse_m2	120	38.5	32.55583	1	100
esse_m3	120	0	446.6576	-1000	1000
esse_m4	120	2533.3	3231.493	1	10000
sel	120	1	0	1	1
mc_neg	60	9.861339	.0171739	9.83221	9.88535
mc_pos	60	9.724487	.0307325	9.674244	9.769611
mcn_neg	60	9.787027	.0135329	9.763389	9.805106
mcn_pos	60	9.665241	.0282369	9.618672	9.706251
mf_neg	60	6.159085	.0148918	6.135935	6.182206

Example for Fuzzy RDD

- 方法三: IV 估计

- ▶ `mcn`, `mf`: 非耐用品和食品支出 (Y)
- ▶ `ess_m`: 已经退休的年数, 负数表示还未到退休年龄
- ▶ `pen`: 退休概率, 内生变量, 依赖于 `ess_m`, 断点处存在不连续跳跃 (X)
- ▶ `elig`: 退休资格虚拟变量, 若 $ess_m \geq 0$, `elig`=1, 否则为 0 (Z)

Example for Fuzzy RDD

- 方法三: IV 估计

```
. ivregress 2sls mcn (pen=elig) esse_m esse_m2 anno1995-anno2004, first robust
First-stage regressions
```

```
Number of obs   =      120
F(      8,      111) =      177.06
Prob > F        =      0.0000
R-squared       =      0.9230
Adj R-squared   =      0.9175
Root MSE       =      0.0965
```

pen	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
esse_m	.0169943	.0031324	5.43	0.000	.0107873	.0232013
esse_m2	-.0005738	.000294	-1.95	0.054	-.0011564	8.82e-06
anno1995	.0230442	.0305492	0.75	0.452	-.0374911	.0835796
anno1998	.0508942	.0336146	1.51	0.133	-.0157153	.1175038
anno2000	.1172279	.0326428	3.59	0.000	.0525441	.1819118
anno2002	.1400281	.0327296	4.28	0.000	.0751722	.204884
anno2004	.1693907	.0333298	5.08	0.000	.1033455	.2354359
elig	.4349947	.0362406	12.00	0.000	.3631815	.5068078
_cons	.2082045	.018284	11.39	0.000	.1719736	.2444355

Example for Fuzzy RDD

- 方法三: IV 估计

```
Instrumental variables (2SLS) regression      Number of obs   =       120
                                             Wald chi2(8)    =       332.96
                                             Prob > chi2     =       0.0000
                                             R-squared       =       0.6223
                                             Root MSE       =       .05749
```

mcn	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
pen	-.0983277	.0544996	-1.80	0.071	-.2051449	.0084895
esse_m	-.0055121	.0025899	-2.13	0.033	-.0105881	-.000436
esse_m2	-.000288	.000145	-1.99	0.047	-.0005722	-3.84e-06
anno1995	.0018884	.0179095	0.11	0.916	-.0332135	.0369903
anno1998	-.0334648	.0181819	-1.84	0.066	-.0691007	.002171
anno2000	.0121598	.019791	0.61	0.539	-.0266299	.0509495
anno2002	.0210096	.0229841	0.91	0.361	-.0240384	.0660575
anno2004	.0843976	.0194665	4.34	0.000	.0462439	.1225512
_cons	9.77691	.0244502	399.87	0.000	9.728988	9.824831

```
Instrumented:  pen
Instruments:  esse_m esse_m2 anno1995 anno1998 anno2000 anno2002 anno2004
              elig
```

Example for Fuzzy RDD

- 方法三: IV 估计

```
. ivregress 2sls mf (pen=elig) esse_m esse_m2 anno1995-anno2004, first robust
First-stage regressions
```

```
Number of obs   =      120
F(      8,      111) =      177.06
Prob > F        =      0.0000
R-squared       =      0.9230
Adj R-squared   =      0.9175
Root MSE       =      0.0965
```

pen	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
esse_m	.0169943	.0031324	5.43	0.000	.0107873	.0232013
esse_m2	-.0005738	.000294	-1.95	0.054	-.0011564	8.82e-06
anno1995	.0230442	.0305492	0.75	0.452	-.0374911	.0835796
anno1998	.0508942	.0336146	1.51	0.133	-.0157153	.1175038
anno2000	.1172279	.0326428	3.59	0.000	.0525441	.1819118
anno2002	.1400281	.0327296	4.28	0.000	.0751722	.204884
anno2004	.1693907	.0333298	5.08	0.000	.1033455	.2354359
elig	.4349947	.0362406	12.00	0.000	.3631815	.5068078
_cons	.2082045	.018284	11.39	0.000	.1719736	.2444355

Example for Fuzzy RDD

- 方法三: IV 估计

```
Instrumental variables (2SLS) regression      Number of obs   =       120
                                             Wald chi2(8)    =       325.13
                                             Prob > chi2     =       0.0000
                                             R-squared      =       0.7124
                                             Root MSE      =       .05365
```

mf	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
pen	-.1409689	.0523376	-2.69	0.007	-.2435487	-.0383891
esse_m	-.0027591	.0024779	-1.11	0.266	-.0076157	.0020975
esse_m2	-.0000821	.0001355	-0.61	0.545	-.0003478	.0001835
anno1995	-.0675309	.0144149	-4.68	0.000	-.0957835	-.0392782
anno1998	-.1365082	.0154907	-8.81	0.000	-.1668694	-.106147
anno2000	-.13855	.0177911	-7.79	0.000	-.1734199	-.1036801
anno2002	-.1420846	.017727	-8.02	0.000	-.176829	-.1073403
anno2004	-.1042948	.0179742	-5.80	0.000	-.1395236	-.0690661
_cons	6.160132	.0223924	275.10	0.000	6.116244	6.20402

```
Instrumented: pen
Instruments: esse_m esse_m2 anno1995 anno1998 anno2000 anno2002 anno2004
             elig
```